

# LEGGE DEI GRANDI NUMERI - LIMITE CENTRALE

Note Title

06/11/2017

$$C = \begin{pmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{pmatrix}$$

$\sigma_1^2, \sigma_2^2 > 0$  e  $C$  definita positiva  
cioè  $\det C > 0, \text{tr} C > 0$

$$f(x, y) = \frac{1}{2\pi\sqrt{\det C}} \exp\left(-\frac{1}{2}(x, y)C^{-1}\begin{pmatrix} x \\ y \end{pmatrix}\right)$$

Se  $(X, Y)$  è una v.e. in  $\mathbb{R}^2$  i.c.  $\mathbb{P}_{X, Y} = f(x, y) dx dy$ ,

allora  $\mathbb{P}_X = N(0, \sigma_1^2)$

$\mathbb{P}_Y = N(0, \sigma_2^2)$

$$f(x, y) = \frac{1}{2\pi\sqrt{\det C}} \exp\left(\frac{-x^2}{2\sigma_1^2}\right) \exp\left(\frac{-(\sigma_2 y - \frac{\sigma_{12}}{\sigma_1} x)^2}{2\det C}\right)$$

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = E[XY] =$$

$$= \int_{\mathbb{R}^2} xy f(x, y) dx dy =$$

$$= \frac{1}{2\pi\sqrt{\det C}} \int_{\mathbb{R}^2} xy \exp\left(\frac{-x^2}{2\sigma_1^2}\right) \exp\left(\frac{-(\sigma_2 y - \frac{\sigma_{12}}{\sigma_1} x)^2}{2\det C}\right) dx dy$$

$$= \frac{1}{2\pi\sqrt{\det C}} \int_{\mathbb{R}} x \exp\left(\frac{-x^2}{2\sigma_1^2}\right) \left( \int_{\mathbb{R}} y \exp\left(\frac{-(\sigma_2 y - \frac{\sigma_{12}}{\sigma_1} x)^2}{2\det C}\right) dy \right) dx$$

$$t = \frac{\sigma_2 y - \frac{\sigma_{12}}{\sigma_1} x}{\sqrt{2\det C}}$$

$$\sigma_2 y - \frac{\sigma_{12}}{\sigma_1} x = t \sqrt{2\det C}$$

$$y = \frac{\sigma_{12}}{\sigma_1^2} x + t \frac{\sqrt{2\det C}}{\sigma_2} \quad dy = \frac{\sqrt{2\det C}}{\sigma_2} dt$$

$$= \frac{1}{2\pi\sqrt{\det C}} \frac{\sqrt{2\det C}}{\sigma_2} \int_{\mathbb{R}} x \exp\left(\frac{-x^2}{2\sigma_1^2}\right) \left( \int_{\mathbb{R}} \left(\frac{\sigma_{12}}{\sigma_1^2} x + t \frac{\sqrt{2\det C}}{\sigma_2}\right) \exp(-t^2) dt \right) dx$$

$$= \frac{1}{\sqrt{2\pi}\sigma_1} \int_{\mathbb{R}} x \exp\left(\frac{-x^2}{2\sigma_1^2}\right) \left(\frac{\sigma_{12}}{\sigma_1^2} x \sqrt{\pi} + 0\right) dx$$

$$= \frac{1}{\sqrt{2\pi}} \frac{b_{12}}{b_1^3} \int_{\mathbb{R}} x^2 \exp\left(\frac{-x^2}{2b_1^2}\right) dx = \frac{b_{12}}{b_1^2} \int_{\mathbb{R}} \frac{x^2}{\sqrt{2\pi b_1^2}} \exp\left(\frac{-x^2}{2b_1^2}\right) dx$$

$$= \frac{b_{12}}{b_1^2} b_1^2 = b_{12}$$

$$f(x|y) = \frac{f(x,y)}{h(y)} = \frac{\frac{1}{2\pi\sqrt{\det C}} \exp\left(\frac{-(b_2^2 x^2 - 2b_{12}xy + b_1^2 y^2)}{2\det C}\right)}{\frac{1}{\sqrt{2\pi b_2^2}} \exp\left(\frac{-y^2}{2b_2^2}\right)}$$

$$= \frac{1}{2\pi\sqrt{\det C}} \sqrt{2\pi b_2^2} \exp\left(\frac{y^2}{2b_2^2} - \frac{b_2^2 x^2 - 2b_{12}xy + b_1^2 y^2}{2\det C}\right)$$

Calculo l' exponente:

$$\frac{1}{2b_2^2 \det C} \left( y^2 (\cancel{b_1^2 b_2^2} - b_{12}^2) - b_2^4 x^2 + 2b_{12} b_2^2 xy - \cancel{b_1^2 b_2^2} y^2 \right)$$

$$= \frac{-1}{2b_2^2 \det C} \left( b_2^4 x^2 - 2b_{12} b_2^2 xy + b_{12}^2 y^2 \right) =$$

$$= \frac{-1}{2b_2^2 \det C} \left( b_2^2 x - b_{12} y \right)^2 = \frac{-1}{2\cancel{b_2^2} \det C} b_2^{4^2} \left( x - \frac{b_{12}}{b_2^2} y \right)^2$$

$$f(x|y) = \frac{1}{\sqrt{2\pi}} \sqrt{\frac{b_2^2}{\det C}} \exp\left(\frac{-(b_2^2 x - b_{12} y)^2}{2b_2^2 \det C}\right) =$$

$$= \frac{1}{\sqrt{2\pi \frac{\det C}{b_2^2}}} \exp\left(\frac{-(x - \frac{b_{12}}{b_2^2} y)^2}{2 \frac{\det C}{b_2^2}}\right)$$

ē la densitē asociate a  $N\left(\frac{b_{12}}{b_2^2} y, \frac{\det C}{b_2^2}\right)$

$$\Rightarrow E[X|Y=y] = \int_{\mathbb{R}} x f(x|y) dx = \frac{b_{12}}{b_2^2} y$$

# LEGGE DEBOLE DEI GRANDI NUMERI

Sia  $\{X_n\}_{n \in \mathbb{N}}$  successione di v.a. su  $(\Omega, \mathcal{E}, \mathbb{P})$ .

Supponiamo che

1) le v.a.  $X_n$  siano a due a due scorrelate  
cioè  $\text{Cov}(X_i, X_j) = 0 \quad i \neq j$

2)  $\mathbb{E}[X_i] = E$  fissa  $\forall i \in \mathbb{N}$

3)  $\exists C^2 > 0$  t.c.  $\text{Var}[X_i] \leq C^2 \quad \forall i \in \mathbb{N}$

Sia  $S_n := \sum_{i=1}^n X_i$

Allora 
$$\int_{\Omega} \left| \frac{S_n}{n}(\omega) - E \right|^2 \mathbb{P}(d\omega) \leq \frac{C^2}{n} \quad \forall n \in \mathbb{N}$$

$$\mathbb{P}\left(\left|\frac{S_n}{n} - E\right| > \delta\right) \leq \frac{C^2}{n\delta^2} \quad \forall n \in \mathbb{N}$$

DIM 
$$S_n := \sum_{i=1}^n X_i$$

$$\mathbb{E}[S_n] = \mathbb{E}\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n \mathbb{E}[X_i] = \sum_{i=1}^n E = nE$$

$$\begin{aligned} \text{Var}[S_n] &= \text{Var}\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n \text{Var}[X_i] + 2 \sum_{i < j} \text{Cov}(X_i, X_j) \\ &\leq \sum_{i=1}^n C^2 = nC^2 \end{aligned}$$

$$\mathbb{E}\left[\frac{S_n}{n}\right] = \frac{1}{n} \mathbb{E}[S_n] = \frac{1}{n} nE = E$$

$$\text{Var}\left[\frac{S_n}{n}\right] = \frac{1}{n^2} \text{Var}[S_n] \leq \frac{1}{n^2} nC^2 = \frac{C^2}{n}$$

$$\int_{\Omega} \left| \frac{S_n}{n}(\omega) - E \right|^2 \mathbb{P}(d\omega) = \mathbb{E}\left[\left(\frac{S_n}{n} - E\right)^2\right] = \text{Var}\left[\frac{S_n}{n}\right] \leq \frac{C^2}{n}$$

$$\mathbb{P}\left(\left|\frac{S_n}{n} - E\right| > \delta\right) \leq \frac{1}{\delta^2} \text{Var}\left[\frac{S_n}{n}\right] \leq \frac{C^2}{\delta^2 n}$$

N.B. Con le ipotesi fatte vale la LEGGE FORTE DEI

# GRANDI NUMERI

$$\lim_{n \rightarrow \infty} \frac{S_n}{n}(\omega) = E \quad \mathbb{P}\text{-q.c.}$$

$$\mathbb{R}^N \quad Q = [0, 1]^N \quad f \in C^0(Q) \quad \int_{\mathbb{R}^N} f(x) dx$$

$(\mathbb{R}^N, \mathcal{B}(\mathbb{R}^N), \mu)$  distribuzione di probabilità

$f: \mathbb{R}^N \rightarrow \mathbb{R}$  limitata  $|f(x)| \leq M \quad \forall x \in \mathbb{R}^N$

$$\bar{f} = \int_{\mathbb{R}^N} f(x) \mu(dx) \quad \int_{\mathbb{R}^N} |f(x)| \mu(dx) \leq M \mu(\mathbb{R}^N) = M \cdot 1 = M$$

Sia  $\{X_n\}_{n \in \mathbb{N}}$  una successione di v.e. su  $(\Omega, \mathcal{E}, \mathbb{P})$

$\mathbb{P}_{X_n} = \mu$  e le  $X_n$  a due a due scorrelate

$$Y_n = f(X_n)$$

$$E[|Y_n|] = \int_{\Omega} |f(X_n(\omega))| \mathbb{P}(d\omega) = \int_{\mathbb{R}^N} |f(x)| \mu(dx) \leq M \mu(\mathbb{R}^N) = \mu \cdot 1 = \mu$$

$$E[Y_n] = \int_{\Omega} f(X_n(\omega)) \mathbb{P}(d\omega) = \int_{\mathbb{R}^N} f(x) \mu(dx) = \bar{f}$$

$$\text{Var}[Y_n] = \int_{\Omega} (Y_n - \bar{f})^2(\omega) \mathbb{P}(d\omega) =$$

$$= \int_{\mathbb{R}^N} (f(x) - \bar{f})^2 \mu(dx) \leq 4M^2 \int_{\mathbb{R}^N} \mu(dx) = 4M^2 \cdot 1 = 4M^2$$

$$|f(x) - \bar{f}| \leq |f(x)| + |\bar{f}| \leq M + M = 2M$$

Applico la legge debole dei grandi numeri alla successione  $Y_n$ ,  $S_n = \sum_{i=1}^n Y_i = \sum_{i=1}^n f(X_i)$

$$\mathbb{P}\left(\left|\frac{1}{n} \sum_{i=1}^n f(X_i) - \bar{f}\right| > \delta\right) \leq \frac{4M^2}{\delta^2 n} \quad \forall n \in \mathbb{N} \quad \forall \delta > 0$$

Supponiamo di voler valutare  $\bar{f}$  con un errore del 5%

$\delta = 5\% = \frac{5}{100} = 5 \cdot 10^{-3}$  e con probabilità di errore al 95% di essere pari al 95%

$$P\left(\left|\frac{1}{n} \sum_{i=1}^n f(X_i) - \bar{f}\right| > 5 \cdot 10^{-3}\right) < 5\% = 5 \cdot 10^{-2}$$

Per questo avere  $\frac{4M^2}{5^2 n} < 5 \cdot 10^{-2}$

$$\frac{4M^2}{(5 \cdot 10^{-3})^2 n} < 5 \cdot 10^{-2}$$

$$n > \frac{4M^2}{25 \cdot 10^{-6}} \cdot \frac{1}{5 \cdot 10^{-2}}$$

$$n > \frac{4M^2 \cdot 10^8}{125}$$

$$x_n = X_n(\omega)$$

$$\frac{1}{n} \sum_{i=1}^n f(x_i)$$

## TEOREMA CENTRALE DEL LIMITE

Sia  $\{X_n\}_{n \in \mathbb{N}}$  successione di v.o. su  $(\Omega, \mathcal{F}, P)$  che siano i.i.d.,  $E[X_n] = E$  finito,  $\text{Var}[X_n] = \sigma^2 > 0$  finite

Per  $n \geq 1$   $S_n := \sum_{k=1}^n X_k$

Allora  $\forall t \in \mathbb{R}$   $P\left(\frac{S_n - nE}{\sigma\sqrt{n}} \leq t\right) \rightarrow \Phi(t)$  per  $n \rightarrow \infty$

Inoltre

$$\lim_{n \rightarrow \infty} \sup_{t \in \mathbb{R}} \left| P\left(\frac{S_n - nE}{\sigma\sqrt{n}} \leq t\right) - \Phi(t) \right| = 0$$

o, equivalentemente

legge di  $\frac{S_n - nE}{\sigma\sqrt{n}}$

$$\lim_{n \rightarrow \infty} \sup_{t \in \mathbb{R}} \left| P(S_n \leq t) - \Phi\left(\frac{t - nE}{\sigma\sqrt{n}}\right) \right| = 0$$

# TEOREMA DI BERRY-ESSEEN

Sia  $\{X_n\}_{n \in \mathbb{N}}$  successione di v.a. i.i.d.

$$E[X_n] = 0$$

$$\text{Var}[X_n] = \sigma^2$$

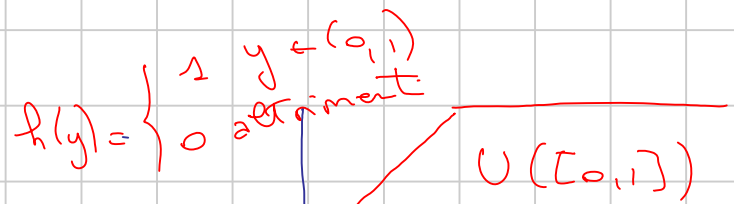
$$E[|X_n|^3] = \gamma < +\infty$$

Sia  $S_n := \sum_{k=1}^n X_k$

$$\text{Allora } \left| P\left(\frac{S_n}{\sigma\sqrt{n}} \leq t\right) - \Phi(t) \right| \leq \frac{C}{\sqrt{n}} \quad \forall t \in \mathbb{R}$$

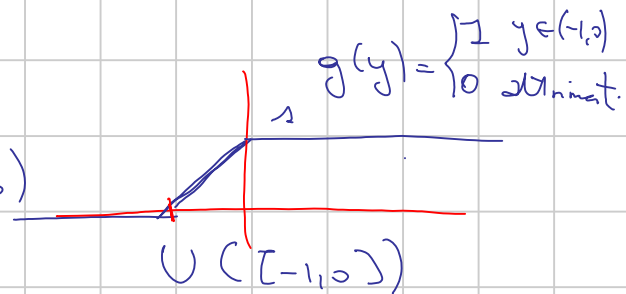
$$C = \frac{0.8 \gamma}{\sigma^3}$$

## ES 1 FOGUOS



$$P_X = B(p)$$

$$P(Y \leq t | X=1) = \begin{cases} 0 & t < 0 \\ t & t \in [0, 1) \\ 1 & t \geq 1 \end{cases}$$



$$P(Y \leq t | X=0) = \begin{cases} 0 & t < -1 \\ t+1 & t \in [-1, 0) \\ 1 & t \geq 0 \end{cases}$$

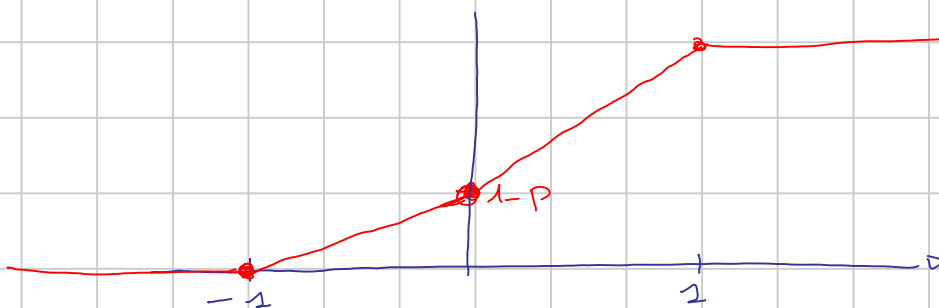
$$P(Y \leq t) = P(Y \leq t | X=1)P(X=1) + P(Y \leq t | X=0)P(X=0)$$

$$t < -1 \quad P(Y \leq t) = 0$$

$$t \in [-1, 0) \quad P(Y \leq t) = (t+1)(1-p)$$

$$t \in [0, 1) \quad P(Y \leq t) = tp + 1 \cdot (1-p)$$

$$t \geq 1 \quad P(Y \leq t) = 1 \cdot p + 1 \cdot (1-p) = 1$$



$$= \int_{-\infty}^t f(y) dy$$

$$f(y) = \begin{cases} 0 & y < -1 \vee y > 1 \\ 1-p & -1 \leq y < 0 \\ p & 0 \leq y < 1 \end{cases}$$

$$\begin{aligned} \mathbb{E}[Y] &= \int_{\mathbb{R}} y f(y) dy = \int_{-1}^0 y(1-p) dy + \int_0^1 y p dy \\ &= \frac{1-p}{2} y^2 \Big|_{y=-1}^{y=0} + \frac{p}{2} y^2 \Big|_{y=0}^{y=1} = \frac{-(1-p)}{2} + \frac{p}{2} = \frac{2p-1}{2} \end{aligned}$$

$$\begin{aligned} \mathbb{E}[Y^2] &= \int_{\mathbb{R}} y^2 f(y) dy = \int_{-1}^0 y^2(1-p) dy + \int_0^1 y^2 p dy \\ &= \frac{1-p}{3} y^3 \Big|_{y=-1}^{y=0} + \frac{p}{3} y^3 \Big|_{y=0}^{y=1} = \frac{1-p}{3} + \frac{p}{3} = \frac{1}{3} \end{aligned}$$

$$\text{Var}[Y] = \frac{1}{3} - \left(\frac{2p-1}{2}\right)^2 = \frac{1}{3} - \frac{4p^2 - 4p + 1}{4} = \frac{1}{12} + p - p^2$$

$$\mathbb{E}[Y|X=t] = \varphi(t)$$

$$\varphi(0) \quad \varphi(1)$$

$$\mathbb{E}[Y|X=0] = \frac{1}{\mathbb{P}(X=0)} \int_{\Omega} Y(\omega) \mathbb{1}_{\{X=0\}}(\omega) \mathbb{P}(d\omega) =$$

$$= \int_{\Omega} Y(\omega) \mathbb{P}_{\{X=0\}}(d\omega) = \int_{\mathbb{R}} y g(y) dy \quad g(y) = \begin{cases} +1 & y \in (-1, 0) \\ 0 & \text{otherwise} \end{cases}$$

$$= \int_{-1}^0 y \cdot 1 dy = \frac{y^2}{2} \Big|_{y=-1}^{y=0} = -\frac{1}{2}$$

$$\mathbb{E}[Y|X=1] = \frac{1}{\mathbb{P}(X=1)} \int_{\Omega} Y(\omega) \mathbb{1}_{\{X=1\}}(\omega) \mathbb{P}(d\omega)$$

$$= \int_{\Omega} Y(\omega) \mathbb{P}_{\{X=1\}}(d\omega) = \int_{\mathbb{R}} y h(y) dy = \int_0^1 y \cdot 1 dy = \frac{y^2}{2} \Big|_{y=0}^{y=1}$$

$$= \frac{1}{2}$$

$$\begin{aligned} E[Y|X](\omega) &= E[Y|X=0] \mathbb{1}_{\{X=0\}}(\omega) + \\ &\quad + E[Y|X=1] \mathbb{1}_{\{X=1\}}(\omega) \\ &= -\frac{1}{2} \mathbb{1}_{\{X=0\}}(\omega) + \frac{1}{2} \mathbb{1}_{\{X=1\}}(\omega) \end{aligned}$$