

DISTRIBUZIONI GANNA. SOMME in V.A. INDIPENDENTI

Note Title

30/10/2017

Se X e Y sono v.e. indipendenti, entrambe con distribuzione A.C.

$$P_X = f(x) dx$$

$$P_Y = g(y) dy$$

Allora la v.e. bidimensionale (X, Y) ha distribuzione A.C. con densità $h(x, y) = f(x)g(y)$

$$P_{X, Y} = f(x)g(y) dx dy$$

(X, Y) v.e. bidimensionale con distribuzione A.C. e densità

$$f(x, y) \quad P_{X, Y} = f(x, y) dx dy$$

$$A \in \mathcal{B}(\mathbb{R}) \quad P_X(A) = P(X \in A) = P(X \in A, Y \in \mathbb{R}) =$$

$$= P((X, Y) \in A \times \mathbb{R}) = \int_{A \times \mathbb{R}} f(x, y) dx dy =$$

$$= \int_A \left(\int_{\mathbb{R}} f(x, y) dy \right) dx$$

$$g(x)$$

$$\forall A \in \mathcal{B}(\mathbb{R}) \quad P_X(A) = \int_A g(x) dx \quad P_X = g(x) dx$$

$$g(x) = \int_{\mathbb{R}} f(x, y) dy$$

$$P_Y = h(y) dy \quad h(y) = \int_{\mathbb{R}} f(x, y) dx$$

X e Y v.e. con distribuzioni di Poisson, indipendenti.

$$P_X = P(\lambda)$$

$$P_Y = P(\mu)$$

$$P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

$$P(Y=k) = e^{-\mu} \frac{\mu^k}{k!}$$

$\forall k = 0, 1, 2, \dots$

$$Z = X + Y$$

$$\forall k \in \mathbb{Z} \quad \mathbb{P}(Z=k) = \sum_{j=-\infty}^{+\infty} \mathbb{P}(X=j) \mathbb{P}(Y=k-j)$$

$$\mathbb{P}(Z=k) = 0 \quad \forall k = -1, -2, \dots$$

$$\begin{aligned} k \geq 0 \quad \mathbb{P}(Z=k) &= \sum_{j=0}^k \mathbb{P}(X=j) \mathbb{P}(Y=k-j) && k-j \geq 0 \quad j \leq k \\ &= \sum_{j=0}^k e^{-\lambda} \frac{\lambda^j}{j!} e^{-\mu} \frac{\mu^{k-j}}{(k-j)!} \frac{k!}{k!} \\ &= e^{-\lambda} e^{-\mu} \frac{1}{k!} \sum_{j=0}^k \binom{k}{j} \lambda^j \mu^{k-j} = (e^{-\lambda} e^{-\mu}) \sum_{j=0}^k \binom{k}{j} \lambda^j \mu^{k-j} \\ &= e^{-(\lambda+\mu)} \frac{1}{k!} (\lambda+\mu)^k = \mathbb{P}_Z = \mathbb{P}(\lambda+\mu) \end{aligned}$$

X e Y v.e. con distribuzioni esponenziale e independent.

$$\mathbb{P}_X = \exp(-\lambda) \quad \mathbb{P}_Y = \exp(-\mu)$$

$$f(x) = \begin{cases} 0 & x \leq 0 \\ \lambda e^{-\lambda x} & x > 0 \end{cases} \quad g(x) = \begin{cases} 0 & x \leq 0 \\ \mu e^{-\mu x} & x > 0 \end{cases}$$

$$\mathbb{P}_Z \quad Z = X+Y \quad \mathbb{P}_Z = h(x) dx \quad h(x) = \int_{\mathbb{R}} f(y) g(x-y) dy$$

$$\begin{cases} y > 0 \\ x-y > 0 \end{cases} \quad 0 < y < x \quad x \leq 0 \quad h(x) = 0 \\ x > 0 \quad h(x) = \int_0^x \lambda e^{-\lambda y} \mu e^{-\mu(x-y)} dy$$

$$x > 0 \quad h(x) = \lambda \mu \int_0^x e^{-\mu x} e^{(\mu-\lambda)y} dy = \lambda \mu e^{-\mu x} \int_0^x e^{(\mu-\lambda)y} dy$$

$$\mu = \lambda \quad h(x) = \lambda^2 x e^{-\lambda x}$$

$$\mu \neq \lambda \quad h(x) = \lambda \mu e^{-\mu x} \frac{e^{(\mu-\lambda)y}}{\mu-\lambda} \Big|_{y=0}^{y=x} = \frac{\lambda \mu}{\mu-\lambda} e^{-\mu x} \left(\frac{e^{(\mu-\lambda)x}}{e^0} - 1 \right)$$

$$= \frac{\lambda^\mu}{\mu - \lambda} \left(e^{-\lambda x} - e^{-\mu x} \right)$$

$$P_X = N(0, \sigma^2) \quad P_Y = N(0, \delta^2) \quad X, Y \text{ independent.}$$

$$P_X = f(x) dx \quad P_Y = g(x) dx$$

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(\frac{-x^2}{2\sigma^2}\right) \quad g(x) = \frac{1}{\sqrt{2\pi\delta^2}} \exp\left(\frac{-x^2}{2\delta^2}\right)$$

$$Z := X + Y \quad P_Z = h(x) dx \quad h(x) = \int_{\mathbb{R}} f(y) g(x-y) dy \quad x \in \mathbb{R}$$

$$h(x) = \frac{1}{2\pi\sigma\delta} \int_{\mathbb{R}} \exp\left(\frac{-y^2}{2\sigma^2}\right) \exp\left(\frac{-(x-y)^2}{2\delta^2}\right) dy$$

$$= \frac{1}{2\pi\sigma\delta} \int_{\mathbb{R}} \exp\left(-\frac{1}{2} \left\{ \frac{y^2}{\sigma^2} + \frac{x^2 - 2xy + y^2}{\delta^2} \right\}\right) dy$$

Coeff d: y^2 $\frac{1}{\sigma^2} + \frac{1}{\delta^2} = \frac{\delta^2 + \sigma^2}{\delta^2\sigma^2}$

coeff d: x^2 $\frac{1}{\delta^2}$

Coeff d: $-2xy$ $\frac{1}{\delta^2}$

$$(ay - bx)^2 + cx^2 =$$

$$a^2y^2 - 2abxy + b^2x^2 + cx^2$$

$$\left(y \frac{\sqrt{\delta^2 + \sigma^2}}{\sigma\delta} - 2y \frac{\sqrt{\delta^2 + \sigma^2}}{\sigma\delta} \times \frac{1}{\delta^2} \frac{\sigma\delta}{\sqrt{\delta^2 + \sigma^2}} \right)^2$$

$$+ x^2 \frac{\sigma^2}{\delta^2(\delta^2 + \sigma^2)} - x^2 \frac{\sigma^2}{\delta^2(\delta^2 + \sigma^2)} + \frac{x^2}{\delta^2}$$

$$= \left(y \frac{\sqrt{\delta^2 + \sigma^2}}{\sigma\delta} - x \frac{\sigma}{\delta\sqrt{\delta^2 + \sigma^2}} \right)^2 + \frac{x^2 (\cancel{\delta^2 + \sigma^2} - \sigma^2)}{\cancel{\delta^2} (\delta^2 + \sigma^2)} =$$

$$= \left(y \frac{\sqrt{\delta^2 + \sigma^2}}{\sigma\delta} - x \frac{\sigma}{\delta\sqrt{\delta^2 + \sigma^2}} \right)^2 + \frac{x^2}{\delta^2 + \sigma^2}$$

$$h(x) = \frac{1}{2\pi\sigma\delta} \int_{\mathbb{R}} \exp\left(-\frac{1}{2} \left\{ \left(y \frac{\sqrt{\delta^2 + \sigma^2}}{\sigma\delta} - x \frac{\sigma}{\delta\sqrt{\delta^2 + \sigma^2}} \right)^2 + \frac{x^2}{\delta^2 + \sigma^2} \right\}\right) dy$$

$$= \frac{1}{2\pi\sigma\delta} \exp\left(\frac{-x^2}{2(\sigma^2+\delta^2)}\right) \int_{\mathbb{R}} \exp\left(-\frac{1}{2}\left(y\frac{\sqrt{\sigma^2+\delta^2}}{\sigma\delta} - x\frac{\sigma}{\delta\sqrt{\sigma^2+\delta^2}}\right)^2\right) dy$$

$$\int_{\mathbb{R}} e^{-t^2} dt = \sqrt{\pi}$$

$$t = \frac{1}{\sqrt{2}} \left(y \frac{\sqrt{\sigma^2+\delta^2}}{\sigma\delta} - x \frac{\sigma}{\delta\sqrt{\sigma^2+\delta^2}} \right)$$

$$dt = \frac{\sqrt{\sigma^2+\delta^2}}{\sigma\delta\sqrt{2}} dy$$

$$h(x) = \frac{1}{\sqrt{2\pi}\sigma\delta} \exp\left(\frac{-x^2}{2(\sigma^2+\delta^2)}\right) \int_{\mathbb{R}} \exp(-t^2) \frac{\cancel{\sigma\delta}\sqrt{2}}{\sqrt{\sigma^2+\delta^2}} dt$$

$$= \frac{1}{\sqrt{2(\sigma^2+\delta^2)}\pi} \exp\left(\frac{-x^2}{2(\sigma^2+\delta^2)}\right) \quad \forall x \in \mathbb{R}$$

$$\Rightarrow \mathbb{P}_{X+Y} = N(0, \sigma^2 + \delta^2)$$

$$\sigma^2 = \delta^2 \quad \mathbb{P}_{X+Y} = N(0, 2\sigma^2)$$

$$X_1, X_2, X_3 \text{ v.e. i.i.d. } \mathbb{P}_{X_i} = N(0, \sigma^2)$$

$$\mathbb{P}_{X_1+X_2} = N(0, 2\sigma^2)$$

$$\mathbb{P}_{\underbrace{(X_1+X_2)}_{N(0, 2\sigma^2)} + X_3}_{N(0, \sigma^2)} = N(0, 3\sigma^2)$$

$$X_1, \dots, X_n \text{ v.e. i.i.d. } \mathbb{P}_{X_i} = N(0, \sigma^2)$$

$$S_n := \sum_{i=1}^n X_i \Rightarrow \mathbb{P}_{S_n} = N(0, n\sigma^2)$$

$$\tilde{X}_n := \frac{S_n}{n}$$

$$\mathbb{E}[\tilde{X}_n] = \mathbb{E}\left[\frac{S_n}{n}\right] = \frac{1}{n} \mathbb{E}[S_n] = 0$$

$$\text{Var}[\tilde{X}_n] = \text{Var}\left[\frac{S_n}{n}\right] = \frac{1}{n^2} \text{Var}[S_n] = \frac{1}{n^2} n\sigma^2 = \frac{\sigma^2}{n}$$

$$\Rightarrow \mathbb{P}_{\tilde{X}_n} = N\left(0, \frac{\sigma^2}{n}\right)$$

$$P_X = \exp(-\lambda) \quad P_Y = \exp(-\mu) \quad X, Y \text{ independent.}$$

$$Z := \min(X, Y)$$

$$Z = f_0(X, Y) \quad f: (x, y) \in \mathbb{R}^2 \mapsto \min(x, y) \in \mathbb{R}$$

$\psi: \mathbb{R} \rightarrow \mathbb{R}$ function of Borel nonnegative

$$\int_{\mathbb{R}} \psi(t) P_Z(dt) = \int_{\mathbb{R}} \psi(t) P_{f_0(X, Y)}(dt) = \int_{\mathbb{R}^2} \psi(f(x, y)) P_{X, Y}(dx dy)$$

$$= \int_{\mathbb{R}} \left(\int_{\mathbb{R}} \psi(\min(x, y)) P_Y(dy) \right) P_X(dx)$$

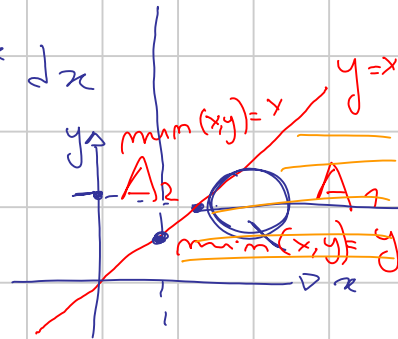
$$P_X = f(x) dx \quad f(x) = \begin{cases} 0 & x \leq 0 \\ \lambda e^{-\lambda x} & x > 0 \end{cases}$$

$$P_Y = g(y) dy \quad g(y) = \begin{cases} 0 & y \leq 0 \\ \mu e^{-\mu y} & y > 0 \end{cases}$$

$$= \int_{\mathbb{R}} \left(\int_{\mathbb{R}} \psi(\min(x, y)) g(y) dy \right) f(x) dx$$

$$\int_0^{+\infty} \left(\int_0^{+\infty} \psi(\min(x, y)) \mu e^{-\mu y} dy \right) \lambda e^{-\lambda x} dx$$

$$= \iint_Q \psi(\min(x, y)) \mu \lambda e^{-\lambda x} e^{-\mu y} dx dy$$



$$= \iint_{A_1} \psi(y) \mu \lambda e^{-\lambda x} e^{-\mu y} dx dy + \iint_{A_2} \psi(x) \mu \lambda e^{-\lambda x} e^{-\mu y} dx dy$$

$$A_1: y \in (0, +\infty) \quad x \in (y, +\infty) \quad A_2: x \in (0, +\infty) \quad y \in (x, +\infty)$$

$$= \int_0^{+\infty} \left(\int_y^{+\infty} \psi(y) \lambda \mu e^{-\lambda x} e^{-\mu y} dx \right) dy + \int_0^{+\infty} \left(\int_x^{+\infty} \psi(x) \lambda \mu e^{-\lambda x} e^{-\mu y} dy \right) dx$$

$$= \int_0^{+\infty} \psi(y) \mu e^{-\mu y} \left(\int_y^{+\infty} \lambda e^{-\lambda x} dx \right) dy + \int_0^{+\infty} \psi(x) \lambda e^{-\lambda x} \left(\int_x^{+\infty} e^{-\mu y} dy \right) dx$$

$$\begin{aligned}
&= \int_0^{+\infty} \varphi(y) \mu e^{-\mu y} \left(-e^{-\lambda x} \right) \Big|_{x=y}^{x=+\infty} dy + \int_0^{+\infty} \varphi(x) \lambda e^{-\lambda x} \left(-e^{-\mu y} \right) \Big|_{y=x}^{y=+\infty} dx \\
&= \int_0^{+\infty} \varphi(y) \mu e^{-\mu y} (0 - (-e^{-\lambda y})) dy + \int_0^{+\infty} \varphi(x) \lambda e^{-\lambda x} (0 - (-e^{-\mu x})) dx \\
&= \int_0^{+\infty} \varphi(x) \mu e^{-\mu x} e^{-\lambda x} dx + \int_0^{+\infty} \varphi(x) \lambda e^{-\lambda x} e^{-\mu x} dx \\
&= \int_0^{+\infty} \varphi(x) (\mu + \lambda) e^{-(\mu + \lambda)x} dx = \int_{\mathbb{R}} \varphi(x) g(x) dx
\end{aligned}$$

$$g(x) = \begin{cases} 0 & x \leq 0 \\ (\mu + \lambda) e^{-(\mu + \lambda)x} & x > 0 \end{cases}$$

$$\mathbb{P}_{\min(x,y)} = \int_{\mathbb{R}} g(x) dx \qquad \mathbb{P}_{\min(x,y)} = \exp(-(\lambda + \mu))$$

DISTRIBUZIONI GAMMA

FUNZIONE GAMMA DI EULERO

$$\alpha > 0 \qquad \int_0^{+\infty} x^{\alpha-1} e^{-x} dx \qquad \alpha > 0 \quad \alpha - 1 > -1$$

$$\int_{\varepsilon}^1 x^p dx = \begin{cases} \log(x) \Big|_{x=\varepsilon}^{x=1} & p = -1 \\ \frac{1}{p+1} x^{p+1} \Big|_{x=\varepsilon}^{x=1} & p \neq -1 \end{cases} \qquad \int_0^1 x^p dx$$

$$= \begin{cases} 0 - \log(\varepsilon) & p = -1 \\ \frac{1}{p+1} (1 - \varepsilon^{p+1}) & p \neq -1 \end{cases} \xrightarrow{\varepsilon \rightarrow 0^+} \begin{cases} 1 - \infty & p = -1 \\ \frac{1}{p+1} & p > -1 \\ +\infty & p < -1 \end{cases}$$

$\forall \alpha > 0$ $\int_0^{+\infty} x^{\alpha-1} e^{-x} dx$ è un integrale convergente

Considero la funzione $\Gamma: \alpha \in (0, +\infty) \mapsto \int_0^{+\infty} x^{\alpha-1} e^{-x} dx$

La funzione Γ si chiama Γ di EULERO

$$\Gamma(1) = \int_0^{+\infty} x^0 e^{-x} dx = \int_0^{+\infty} e^{-x} dx = -e^{-x} \Big|_{x=0}^{x \rightarrow +\infty} = 1$$

$$\Gamma(n+1) = \int_0^{+\infty} x^{n+1-1} e^{-x} dx = \int_0^{+\infty} x^n e^{-x} dx =$$

$$x^n (-e^{-x}) \Big|_{x=0}^{x \rightarrow +\infty} + \int_0^{+\infty} n x^{n-1} e^{-x} dx = n \int_0^{+\infty} x^{n-1} e^{-x} dx$$

$$\boxed{\Gamma(n+1) = n \Gamma(n)}$$

$$\Gamma(1) = 1$$

$$\Gamma(2) = 1 \cdot \Gamma(1) = 1 \cdot 1 = 1$$

$$\Gamma(3) = 2 \Gamma(2) = 2 \cdot 1$$

$$\Gamma(n) = (n-1)!$$

$$\Gamma(4) = 3 \Gamma(3) = 3 \cdot 2 \cdot 1$$

$$\Gamma(5) = 4 \Gamma(4) = 4 \cdot 3 \cdot 2 \cdot 1$$

Per λ, α parametr. reali positivi chiamo DISTRIBUZIONE
GAMMA DI PARAMETRI α e λ , la distribuzione $\Gamma(\alpha, \lambda)$
r.c. associata alle unità

$$f(x) = \begin{cases} 0 & x \leq 0 \\ \frac{\lambda^\alpha x^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} & x > 0 \end{cases}$$

$$f(x) \geq 0 \quad \forall$$

$$\int_{\mathbb{R}} f(x) dx = \int_0^{+\infty} \frac{\lambda^\alpha x^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} dx = \frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^{+\infty} x^{\alpha-1} e^{-\lambda x} dx$$

$$t = \lambda x \quad x = \frac{t}{\lambda} \quad dx = \frac{1}{\lambda} dt$$

$$\frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^{+\infty} \frac{t^{\alpha-1}}{\lambda^{\alpha-1}} e^{-t} \frac{1}{\lambda} dt = \frac{1}{\Gamma(\alpha)} \Gamma(\alpha) = 1$$

$$F_X = \Gamma(\alpha, \lambda) \quad X \geq 0 \text{ p.c.}$$

$$\begin{aligned} E[X] &= \int_{\mathbb{R}} x f(x) dx = \int_0^{+\infty} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^\alpha e^{-\lambda x} dx && t = \lambda x \\ &= \frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^{+\infty} \frac{t^\alpha}{\lambda^\alpha} e^{-t} \frac{1}{\lambda} dt = \frac{1}{\lambda \Gamma(\alpha)} \int_0^{+\infty} t^{\alpha-1} e^{-t} dt && x = \frac{t}{\lambda} \\ &= \frac{1}{\lambda \Gamma(\alpha)} \Gamma(\alpha+1) = \frac{\alpha \Gamma(\alpha)}{\lambda \Gamma(\alpha)} = \frac{\alpha}{\lambda} \end{aligned}$$

$$\begin{aligned} E[X^2] &= \int_{\mathbb{R}} x^2 f(x) dx = \int_0^{+\infty} \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha+1} e^{-\lambda x} dx && t = \lambda x \\ &= \frac{\lambda^\alpha}{\Gamma(\alpha)} \int_0^{+\infty} \frac{t^{\alpha+1}}{\lambda^{\alpha+1}} e^{-t} \frac{1}{\lambda} dt = \frac{1}{\lambda^2 \Gamma(\alpha)} \int_0^{+\infty} t^{\alpha+1} e^{-t} dt && x = \frac{t}{\lambda} \\ &= \frac{1}{\lambda^2 \Gamma(\alpha)} \Gamma(\alpha+2) = \frac{1}{\lambda^2 \Gamma(\alpha)} (\alpha+1) \Gamma(\alpha+1) = \end{aligned}$$

$$= \frac{\alpha+1}{\lambda^2 \Gamma(\alpha)} \alpha \Gamma(\alpha) = \frac{(\alpha+1)\alpha}{\lambda^2}$$

$$\text{Var}[X] = \frac{(\alpha+1)\alpha}{\lambda^2} - \left(\frac{\alpha}{\lambda}\right)^2 = \frac{\alpha}{\lambda^2} (\alpha+1 - \alpha) = \frac{\alpha}{\lambda^2}$$